# Assessment of participation bias in cohort studies: systematic review and metaregression analysis 

Avaliação do viés de participação em estudos de coorte: uma revisão sistemática e metarregressão Evaluación del sesgo de participación en estudios de cohortes: una revisión sistemática y metarregresión

${ }^{1}$ Escola Nacional de Saúde Pública Sergio Arouca, Fundação Oswaldo Cruz, Rio de Janeiro, Brasil.
2 Instituto Benjamin Constant, Rio de Janeiro, Brasil.
${ }^{3}$ Instituto de Estudos em Saúde Coletiva, Universidade Federal do Rio de Janeiro, Brasil.
4 Programa de Computação Científica, Fundação Oswaldo Cruz, Rio de Janeiro, Brasil.

## Correspondence

S. H. A. Silva Junior

Rua Jaguari 79, Nova Iguaçu, RJ 26010-425, Brasil. sergio.edfisica@gmail.com


#### Abstract

The proportion of non-participation in cohort studies, if associated with both the exposure and the probability of occurrence of the event, can introduce bias in the estimates of interest. The aim of this study is to evaluate the impact of participation and its characteristics in longitudinal studies. A systematic review (MEDLINE, Scopus and Web of Science) for articles describing the proportion of participation in the baseline of cohort studies was performed. Among the 2,964 initially identified, 50 were selected. The average proportion of participation was $64.7 \%$. Using a meta-regression model with mixed effects, only age, year of baseline contact and study region (borderline) were associated with participation. Considering the decrease in participation in recent years, and the cost of cohort studies, it is essential to gather information to assess the potential for non-participation, before committing resources. Finally, journals should require the presentation of this information in the papers.


Selection Bias; Cohort Studies; Epidemiologic Methods
Sérgio Henrique Almeida da Silva Junior 1,2
Simone M. Santos 1
Cláudia Medina Coeli ${ }^{3}$
Marilia Sá Carvalho ${ }^{4}$

## Resumo

A proporção de não-participação em estudos de coorte está associada também à exposição e à probabilidade de ocorrência do evento poder gerar viés nas estimativas de interesse. O objetivo do presente trabalho é realizar uma revisão sistemática e metanálise de artigos que descrevem a participação em estudos de coorte e avaliar as características associadas à participação. Foi realizada uma revisão sistemática (MEDLINE, Scopus e Web of Science), buscando-se artigos que descrevessem a proporção de participação na linha de base de estudos de coorte. De 2.964 artigos inicialmente identificados, foram selecionados 50. Entre esses, a proporção média de participação foi de 64,7\%. Utilizando-se o modelo de metarregressão com efeitos mistos, somente a idade, ano da linha de base e a região do estudo (limítrofe) estiveram associados à participação. Considerando a diminuição na participação em anos mais recentes e o custo dos estudos de coorte, é essencial buscar informações que permitam avaliar o potencial de não-participação antes de comprometer os recursos.

Viés de Seleção; Estudos de Coortes; Métodos Epidemiológicos

## Background

Among observational studies, the advantages of prospective cohort studies are that they are able to estimate incidence measures directly and are less vulnerable to information bias. However, participation refusal at baseline or follow-up can introduce selection bias when simultaneously associated with both exposure and the outcome 1,2 . As a result, the association between exposure and outcome may differ between participants and non-participants.

Morton et al. ${ }^{3}$ observed a tendency for participation in cohort studies to decrease between 1970 and 2003. As the non-participation proportion rises, vulnerability to selection bias tends to increase. Therefore, it is recommended reporting participation proportion in observational studies ${ }^{4}$, designing methodological studies to evaluate the impacts of non-participation and evaluating study characteristics that may influence participation 5 .

To the best of our knowledge, and in spite of its importance, no systematic evaluation of participation in observational cohort studies is available to guide choices and scientific assessment of validity of conclusions. This present study aims to perform a systematic review and meta-regression of papers describing non-participation bias in cohort studies, and evaluate the studies' characteristics associated with participation proportion.

## Methods

We performed a systematic review and meta-regression following the methodology proposed by Higgins \& Green ${ }^{6}$ and PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) criteria ${ }^{7}$.

## Search strategy

We searched MEDLINE, Scopus and Web of Science data bases for papers published between January 1978 and November 2014. The query used for the MEDLINE search strategy was: (cooperation[Title/Abstract/ MESH] or noncooperation[Title/Abstract/ MESH] or non-cooperation[Title/Abstract/ MESH] or participant*[Title/Abstract/ MESH] or nonparticipant*[Title/Abstract/ MESH] or non-participant*[Title/Abstract/ MESH] or compliance[Title/Abstract/MESH] or noncompliance[Title/Abstract/MESH] or non-compliance[Title/Abstract/MESH]) AND bias*[Title/Abstract/MESH] AND (cohort*[Title/ Abstract/MESH] OR prospective [Title/Abstract/ MESH] OR longitudinal [Title/Abstract/MESH]).

For the other data bases, the specific syntaxes corresponding to each base were used.

Article titles and abstracts were evaluated by two reviewers working independently in order to ascertain whether they met the criteria for inclusion in the study. Disagreements were assessed by a third reviewer.

## Eligibility criteria and data extraction

As specific populations and health problems may induce large differences in participation proportions related to theses specificities, we only included population-based cohort studies on adult ( 18 to 75 years old) healthy people. We excluded studies that addressed specific populations (eg. pregnant women, patients with specific ailments), review studies and others (eg. genetic studies, surgery, drug therapies). Figure 1 depicts the review flow chart.

The references identified were stored and processed using the JabRef 2.10 software (http:// jabref.sourceforge.net/). We collected the participation proportion, the general characteristics of the study (year of baseline contact, place, selection strategy and study outcome). We also evaluated the characteristics of the study population including type (general population vs. working population), participation of women and the mean age. The relevant data was extracted reading the full paper.

## Data analysis

A meta-analysis of participation proportion was conducted using mixed-effects models, often called binominal-normal models ${ }^{8}$. Given the heterogeneity of the studies ( $\mathrm{I}^{2}=99.97 \%$; $\tau^{2}$ $=0.54 ; \mathrm{p}<0.001$ ), we investigated the variables associated with the participation proportion, initially by simple meta-regression models. When the value of variance accounted for (VAF) by the model was greater than $5 \%$, the variable was included in the multiple model. VAF indicates the percentage of total heterogeneity that is explained by each moderator. The goodness of fit of the multiple model was evaluated by the likelihood ratio test (LRT).

We analyzed the following variables: year of the baseline contact, participant mean age, proportion of women, selection strategy, population type (general population vs. employees population), study outcome - cardiovascular (baseline category), general health or others (cancer, accident, substance use, incapacity and smoking) - and study region, as divided by United Nations Statistics Division 9 into Continental Europe (baseline category), Northern Europe, USA, and

Others (Asia or Oceania). Spearman correlation coefficient was used to evaluate the relation between the year of the baseline contact and the participation proportion.

The analyses were performed using the metafor ${ }^{10}$ library of R software (The R Foundation for Statistical Computing, Vienna, Austria; http:// www.r-project.org).

## Results

Of the 2,964 original papers initially identified, 50 were selected. Figure 1 summarizes the study selection process

Table 1 describes the objectives, database, analysis and main results of the selected papers. To evaluate participation, 29 ( $58 \%$ ) papers compared participants and non-participants using secondary databases, 15 (30\%) used the information available at baseline, and six ( $12 \%$ ) used some way of contacting the non-respondents with small questionnaires. Logistic regression models were the most used technique to evaluate participation, used in $18(40 \%)$ of the papers. Passive follow-up studies applied survival (7) and Poisson regression models (4), and a few some combination of different techniques. In eight papers the evaluation was based on frequencies comparison, using baseline characteristics and/or questionnaires. Imputation, weighted regression and simulations were applied in four papers to evaluate and propose analytical methods for correcting potential bias.

Table 2 describes of the overall study characteristics and sample characteristics potentially associated with participation proportion. Most of the publications are concentrated in the years from 2005 to 2014, the oldest having been published in 1978. The studies comprised $40(80 \%)$ geographically population-based, while the remainder were of workers (8), students (1) and recruits (1).

Most of the studies were conducted in Northern Europe ( $40 \%$ ). Regarding participant selection, $60 \%$ were random sample, the remainder census-based. The most frequent outcomes were overall health condition in twenty-three ( $46 \%$ ), and cardiovascular health in forteen. Other outcomes included cancer, accident, substance use, incapacity and smoking. Participant mean age was 49.5 years ( $\mathrm{SD}=8.2$ years). Mean participation proportion was $64.7 \%$, and ranged from $32.2 \%$ to $87.3 \%$. Women participation was slightly larger (52.6\%) (Table 2).

A negative correlation was found between study year and participation proportion ( $\rho=$ -0.38 ). Figure 2 shows the downward trend in participation proportion. The dotted line indicates
the linear regression, an annual rate of decrease of $0.66 \%\left(R^{2}=0.1 ; p=0.01\right)$. The continuous line (a smooth spline) indicates a downward trend in participation, since 1985. The diameters of the circles of each study, identified by the number of the study (id) in Table 1, is proportional to the inverse of the corresponding standard errors in the meta-regression. The larger circles are more influential in the meta-regression.

The simple meta-regression showed association only between participation proportion and year of the baseline contact ( $O R=0.97 ; 95 \% \mathrm{CI}$ : $0.95-0.99)$. The multiple meta-regression showed an association between participation proportion, year of the baseline contact ( $\mathrm{OR}=0.97 ; 95 \% \mathrm{CI}$ : $0.95-0.99$ ) and age ( $\mathrm{OR}=0.97$; $95 \% \mathrm{CI}$ : $0.95-1.00$ ) (Table 3). In other words, for one-year increase in the year of the baseline contact of the study we expect a $3 \%$ decrease in the odds of study participation. Likewise, for one-year increase in the mean age of the study participants we expect a $3 \%$ reduction in the odds of study participation.

The analysis shows residual heterogeneity $\tau^{2}=0.41(\mathrm{p}<0.001)$ for the participation proportion, suggesting that $18.1 \%$ of total heterogeneity can be accounted for by including year of the baseline contact and age. The test for residual heterogeneity is significant $(\mathrm{LRT}=42,252.5, \mathrm{df}=33, \mathrm{p}=$ 0.00 ), indicating that other covariates not considered in the model are influencing the participation proportion.

## Discussion

We found a high heterogeneity in participation proportions among the papers evaluating nonparticipation bias. The most referred characteristics described in the systematic reviewed papers were sociodemographic profile, hospitalization and cancer incidence. Mortality was larger among non-participants. However, in the meta-regression performed only year of the baseline contact and age was associated with participation.

Several strategies involving comparison between participants and non-participants have been proposed to evaluate the potential selection bias in cohort studies: questionnaires to non-participants, comparison of participants according to recruitment moment ${ }^{4}$ and passive monitoring of the eligible population using secondary database to assess the outcome ${ }^{11}$, the majority of papers in our study.

The results show a decrease in participation in studies over time. The reasons for this decline are not clear, but social changes, and changes in selection and recruitment and in study designs may influence participation ${ }^{3}$. The decrease in participa-

Figure 1
Flowchart of the search and selection of studies included in the meta-analysis.


Source: Moher et al. 7.
tion may be related particularly to the increasing number of studies in recent decades, as well as the proliferation of political and marketing surveys ${ }^{5}$. In addition, increased requests for biological material in epidemiological studies may influence adherence negatively 3 .

Previous studies have reported the association between young age and participation cohort studies. Contrary to other articles $12,13,14,15$ the proportion of women in the studies showed no association with participation, not even in the simple model. The outcome of the studies was not associ-

Table 1
Characteristics of studies potentially associated with participation.

| Id | Reference | Baseline year | Source population | Study region | Outcome | Selection | Mean age | N | Participation rate (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Studer et al. 37 | 2010 | Recruits | Continental Europe | Other | Sampling | 20 | 5,457 | Overall $=46.2$ |
| 2 | Kaerlev et al. 38 | 2007 | Workers in general | Northern Europe | General health | Sampling | 45 | 4,489 | $\begin{gathered} \text { Overall }=44.7 ; \text { Men = } \\ \text { 79.2; Women }=20.8 \end{gathered}$ |
| 3 | Langley et al. 39 | 2007 | General population | Other | Other | Sampling | 41.4 | 2,856 | $\begin{gathered} \text { Overall = 79.9; Men = } \\ 61.0 ; \text { Women }=39.0 \end{gathered}$ |
| 4 | Alkerwi et al. 40 | 2006 | General population | Continental Europe | Cardiovascular diseases | Sampling | 44.3 | 1,432 | $\begin{gathered} \text { Overall }=32.2 ; \text { Men }= \\ 48.7 ; \text { Women }=51.3 \end{gathered}$ |
| 5 | Langhammer et <br> al. 41 | 2006 | General population | Northern Europe | General health | Census | 53.1 | 50,807 | $\begin{gathered} \text { Overall }=54.1 ; \text { Men }= \\ 45.4 ; \text { Women }=54.6 \end{gathered}$ |
| 6 | Eriksson et al. 42 | 2005 | General population | Northern Europe | Cardiovascular diseases | Census | 47 | 25,173 | $\begin{gathered} \text { Overall }=82.6 ; \text { Men }= \\ 39.4 ; \text { Women }=60.6 \end{gathered}$ |
| 7 | Osler et al. 43 | 2004 | General population | Northern Europe | General health | Census | 51 | 6,292 | Overall $=66.2$ |
| 8 | Buckley et al. 44 | 2003 | General population | Northern Europe | Cardiovascular diseases | Sampling | 63.9 | 493 | $\begin{gathered} \text { Overall = 45.6; Men = } \\ 80.9 ; \text { Women }=19.1 \end{gathered}$ |
| 9 | Schmidt et al. 45 | 2003 | General population | Continental Europe | General health | Sampling | 46.7 | 7,189 | Overall $=64.5$; Men $=$ 45.7; Women $=54.3$ |
| 10 | Martikainen et al. 46 | 2002 | Workers in general | Northern Europe | General health | Sampling | 49.6 | 8,960 | Overall $=67.1$ |
| 11 | Holden et al. 47 | 2001 | General population | Other | General health | Census | 65 | 1,115 | $\begin{gathered} \text { Overall = 42.6; Men = } \\ 49.9 ; \text { Women }=50.1 \end{gathered}$ |
| 12 | Lissner et al. 17 | 2001 | General population | Northern Europe | General health | Sampling | 46.8 | 850 | Overall $=71$ |
| 13 | Stang et al. 16 | 2001 | General population | Continental Europe | Cardiovascular diseases | Sampling | 58.8 | 8,413 | $\begin{gathered} \text { Overall = 53.3; Men = } \\ 54.3 ; \text { Women }=45.7 \end{gathered}$ |
| 14 | Goldberg et al. 20 | 2000 | Electric and gas utility workers | Continental Europe | General health | Census | 45.1 | 20,328 | $\begin{gathered} \text { Overall = 44.1; Men = } \\ 72.9 ; \text { Women }=27.1 \end{gathered}$ |
| 15 | Taylor et al. 48 | 2000 | General population | Other | General health | Sampling | 46 | 6,073 | $\begin{gathered} \text { Overall }=49.6 ; \text { Men = } \\ 48.9 ; \text { Women }=51.1 \end{gathered}$ |
| 16 | Alonso et al. 49 | 1999 | Students | Continental Europe | Cardiovascular diseases | Census | 35.4 | 9,907 | $\begin{gathered} \text { Overall = 87.3; Men = } \\ 40.7 ; \text { Women }=59.3 \end{gathered}$ |
| 17 | Knudsen et al. 19 | 1999 | General population | Northern Europe | Cardiovascular diseases | Census | 48.8 | 18,565 | $\begin{gathered} \text { Overall }=63.2 ; \text { Men }= \\ 51.2 ; \text { Women }=48.8 \end{gathered}$ |
| 18 | Manjer et al. 29 | 1999 | General population | Northern Europe | Other | Census | 52.9 | 28,098 | $\begin{gathered} \text { Overall }=60.5 ; \text { Men }= \\ 39.4 ; \text { Women }=60.6 \end{gathered}$ |
| 19 |  <br> Balzi 24 | 1998 | General population | Continental Europe | Other | Sampling | 61.8 | 1,776 | Overall $=85.8$; Men $=$ 44.3; Women = 55.7 |
| 20 | Bergman et al. 50 | 1998 | General population | Northern Europe | Other | Census | 42.7 | 19,742 | Overall $=52.9$; Men $=$ 44.5; Women $=55.5$ |
| 21 | Petersen et al. 51 | 1998 | General population | Northern Europe | Other | Census | 63.9 | 791 | $\begin{gathered} \text { Overall = 38.4; Men = } \\ 41.8 ; \text { Women }=58.2 \end{gathered}$ |
| 22 | Rao et al. 52 | 1998 | Radiologists | U.S.A. | Other | Census | 50.1 | 90,305 | Overall $=68.4$ |
| 23 | Haring et al. 53 | 1997 | General population | Continental Europe | General health | Sampling | 54.1 | 7,008 | $\begin{gathered} \text { Overall }=47.1 ; \text { Men }= \\ 48.2 ; \text { Women }=51.8 \end{gathered}$ |
| 24 | Van Loon et al. 30 | 1997 | General population | Continental Europe | General health | Sampling | 42.2 | 12,097 | Overall $=56.5$; Men $=$ 44.5; Women $=55.5$ |
| 25 | Drivsholm et al. 54 | 1996 | General population | Northern Europe | Other | Census | 60 | 1,077 | Overall $=64.5$; Men $=$ 46.8; Women = 53.2 |

[^0]Table 1 (continued)

| Id | Reference | Baseline year | Source population | Study region | Outcome | Selection | Mean age | N | Participation rate (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 26 | Jackson et al. 15 | 1996 | General population | U.S.A. | Cardiovascular diseases | Sampling | 54 | 15,800 | Overall $=59.8$; Men $=$ 45.1; Women = 54.9 |
| 27 | Veenstra et al. 28 | 1996 | General population | Continental Europe | Cardiovascular diseases | Sampling | 55.8 | 15,896 | $\begin{gathered} \text { Overall }=51.9 ; \text { Men = } \\ 50.5 ; \text { Women }=49.5 \end{gathered}$ |
| 28 | Young et al. 55 | 1996 | General population | Other | General health | Sampling | 47.5 | 40,395 | Overall $=80.4$ |
| 29 | Caetano et al. 56 | 1995 | General population | U.S.A. | Other | Sampling | 42.2 | 3,106 | $\begin{gathered} \text { Overall = 81.8; Men = } \\ 48.2 ; \text { Women }=51.8 \end{gathered}$ |
| 30 | Garcia et al. 27 | 1994 | General population | Continental Europe | General health | Sampling | 39.1 | 1,438 | Overall $=57.5$; Men $=$ 46.5; Women $=53.5$ |
| 31 | Hara et al. 57 | 1994 | General population | Other | Other | Sampling | 55.5 | 61,447 | Overall $=50.5$; Men $=$ 46.7; Women = 53.8 |
| 32 |  <br> Thoning 32 | 1994 | General population | Northern Europe | General health | Sampling | 45.6 | 18,292 | $\begin{gathered} \text { Overall }=79.2 ; \text { Men }= \\ 48.5 ; \text { Women }=51.5 \end{gathered}$ |
| 33 | Jacobsen et al. 31 | 1993 | General population | U.S.A. | Cardiovascular diseases | Sampling | 60.8 | 963 | Overall $=50.6$; Men $=$ <br> 47.3; Women = 52.7 |
| 34 | Montgomery et al. 58 | 1993 | Pesticide applicators | U.S.A. | General health | Census | 47.3 | 50,764 | $\begin{gathered} \text { Overall = 65.9; Men = } \\ \text { 97.0; Women }=3.0 \end{gathered}$ |
| 35 | Jousilahti et al. 59 | 1992 | General population | Northern Europe | General health | Sampling | 48.1 | 6,051 | Overall $=84.4$; Men $=$ <br> 47.1; Women = 52,9 |
| 36 | May et al. 60 | 1992 | General population | Other | General health | Sampling | 52.2 | 375,815 | $\begin{gathered} \text { Overall = 81.6; Men = } \\ 27.6 ; \text { Women }=72.4 \end{gathered}$ |
| 37 | Batty \& Gale 61 | 1991 | General population | Northern Europe | Cardiovascular diseases | Sampling | 51 | 6,484 | Overall $=70.8$; Men $=$ 44.7; Women $=55.3$ |
| 38 | Dugué et al. 62 | 1990 | General population | Northern Europe | General health | Census | 33.2 | 1,156,671 | Overall $=78.1$ |
| 39 | Hara et al. 22 | 1990 | General population | Other | Cardiovascular diseases | Census | 49.6 | 43,140 | $\begin{gathered} \text { Overall = 79.3; Men = } \\ 48.0 ; \text { Women }=52.0 \end{gathered}$ |
| 40 | Benfante et al. 63 | 1989 | General population | U.S.A. | Cardiovascular diseases | Census | 54.3 | 8,006 | Overall $=71.9$ |
| 41 | Ferrie et al. 23 | 1988 | Office workers | Northern Europe | General health | Census | 46.3 | 10,297 | $\begin{gathered} \text { Overall = 87.1; Men = } \\ 67.0 ; \text { Women }=33.0 \end{gathered}$ |
| 42 | François et al. 64 | 1987 | General population | Continental Europe | Other | Sampling | 43,3 | 1,910 | $\begin{gathered} \text { Overall = 83.1; Men = } \\ 48.9 ; \text { Women }=51.1 \end{gathered}$ |
| 43 | Walker et al. 21 | 1987 | General population | Northern Europe | Cardiovascular diseases | Sampling | 46.4 | 15,364 | Overall $=74.3$ |
| 44 | David et al. 65 | 1986 | General population | Other | Other | Sampling | 48.7 | 2,095 | Overall $=78.0 ;$ Men $=$ <br> 43.8; Women $=56.2$ |
| 45 | Froom et al. 66 | 1985 | Industrial employees | Other | General health | Census | 45 | 5,302 | Overall $=71.6$ |
| 46 | Bopp et al. 67 | 1984 | General population | Continental Europe | Cardiovascular diseases | Census | 47.6 | 10,160 | $\begin{gathered} \text { Overall }=33.9 ; \text { Men }= \\ 49.1 ; \text { Women } 50.9 \end{gathered}$ |
| 47 | Criqui et al. 68 | 1978 | General population | U.S.A. | Cardiovascular diseases | Sampling | 52.5 | 5,052 | $\begin{gathered} \text { Overall }=82.1 ; \text { Men }= \\ 46.0 ; \text { Women }=54.0 \end{gathered}$ |
| 48 | Lindsted et al. 69 | 1976 | General population | U.S.A. | General health | Census | 53 | 39,886 | $\begin{gathered} \text { Overall = 78.0; Men = } \\ 40.8 ; \text { Women }=59.2 \end{gathered}$ |
| 49 | Thygesen et al. 70 | 1976 | General population | Northern Europe | General health | Sampling | 53.1 | 24,464 | $\begin{gathered} \text { Overall = 72.0; Men = } \\ 45.8 ; \text { Women }=54.2 \end{gathered}$ |
| 50 |  <br> Rasmussen 71 | 1974 | Workers in general | Northern Europe | Other | Sampling | 55.1 | 1,404 | Overall $=66.1$ |

Objectives, database, analysis and results of the selected papers

| Id | Reference | Objectives * | Data source | Analysis | Results |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Studer et al. 38 | To evaluate differences in substance use between late and early respondents, non-consenters or silent refusers, and whether converting former non-respondents can reduce non-response bias | Baseline information | Logistic model | Late respondents presented a midway pattern of substance use higher than early respondents, but lower than non-consenters |
| 2 | Kaerlev et al. 39 | To examine bias on the association between occupational stressors and mental health due to non-participation in a prospective cohort | Secondary data | Survival model | Proportions of gender, age, employment status, sick leave and hospitalization for affective disorders were different in respondents and non-respondents, but low participation at baseline was not associated with mental health outcome |
| 3 | Langley et al. 40 | To evaluate factors associated with non-participation in two follow-up contacts of a prospective cohort study of injury outcomes | Baseline information | Poisson model | Non-participation in the closest follow-up contact did not mean non-participation in the next contact; sociodemographic factors were the most important for non-participation |
| 4 | Alkerwi et al. 41 | To evaluate the representativeness of the sample with respect to the population and compare characteristics of participants and non participants | Baseline <br> information | Logistic model | Non-participants were similar to participants in gender and place of residence; younger people were under-represented while adults and elderly were overrepresented; no discriminating health profiles were detected |
| 5 | Langhammer et al. 42 | To study potential participation bias for common symptoms, diseases and socioeconomic status and mortality by participation status | Secondary data, mailed questionnaire. | Negative binomial and survival models | Questionnaire answers indicated higher prevalences of cardiovascular diseases, diabetes mellitus and psychiatric disorders among non-participants; registry data showed higher mortality and lower socioeconomic status among non-participants |
| 6 | Eriksson et al. 43 | To assess selective non-response in population-based cohort study on type 2 diabetes, using the populationbased drug register for the Stockholm Diabetes Prevention Program | Secondary data | Logistic model | At baseline, non-participants and participants were similar. At follow-up, risks were higher among nonparticipants |
| 7 | Osler et al. 44 | To evaluate changes in association measures in early-life aspects and later health outcomes due to non-response in a follow-up survey | Secondary data | Logistic model and comparison of odds ratios between respondents and complete cohort | A low response rate at age 50 years was related to having a single mother at birth, low educational attainment at age 18, and low cognitive function at ages 12 and 18 . The risk of alcohol overuse and tobacco-related diseases was also highest among non-respondents |
| 8 | Buckley et al. 45 | To assess baseline differences in participation in a secondary prevention of ischemic heart disease program | Secondary data | Logistic model | Enrollment was lower for women in general and for men with uncontrolled total cholesterol level |
| 9 | Schmidt et al. 46 | To identify back-pain-related indicators that could predict attrition in longitudinal studies | Baseline information | Logistic model | The best predictors of attrition were age and baseline response behavior. No bias was found in relation to back pain indicators |
| 10 | Martikainen et al. 47 | To estimate impact on social class inequalities in health due to nonresponse | Secondary data | Linear regression model | Higher social class employees and women were more likely to participate, and sickness absence was higher in non-respondents. Social classes differences did not impact sickness absence in participants or nonparticipants |

(continues)

Table 2 (continued)

| Id | Reference | Objectives * | Data source | Analysis | Results |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 11 | Holden et al. 48 | To explore reasons for nonparticipation in a chronic disease management program | Secondary data | Logistic and multinomial model | Reasons for loss-to-follow-up were: refusals - related to older age, female gender and heart failure; untraceable people - younger, single, indigenous; and death - older individuals, male, who had cancer or heart failure |
| 12 | Lissner et al. 18 | To describe 32 years of follow-up of a cohort of women receiving several health examinations | Baseline information, home visits to non-respondents | Linear regression model | Among the $64 \%$ of survivors, non-participants and home visited subjects were similar in regard to anthropometry and blood pressure, and both groups were similar to participants in social indicators |
| 13 | Stang et al. 17 | To compare recruitment strategies and baseline characteristics of participants and non-participants | Sample of the population | Frequencies comparison | Nonparticipants were more often smokers and of lower social class. A regular relationship with a partner was more frequent among participants |
| 14 | Goldberg et al. 21 | To evaluate several variables associated with participation in the French GAZEL cohort | Baseline information | Mixed effects logistic model | Male and older employees in managerial position or retired presented higher response rates. Smoking and alcohol drinking predicted lower participation. Health problems were strong predictors of attrition |
| 15 | Taylor et al. 49 | To analyze the association between health-related and socio-demographic indicators and participation in a biomedical cohort study | Sample of the population | Frequencies comparison | Cohort participants were similar to the source population, except for alcohol consumption, which, at an intermediate to high risk level was more frequent among participants |
| 16 | Alonso et al. 50 | To evaluate potential predictors of retention in a cohort study and selection bias effect in rate ratio estimates due to loss-to-follow-up | Baseline <br> information | Inverse <br> probability weight logistic model | Several variables (age, smoking, marital status, obesity, past vehicle injury and self-reported history of cardiovascular disease) were associated with the probability of attrition. Obesity, when adjusted for confounding, was similarly associated with hypertension in models with and without inverse probability weight |
| 17 | Knudsen et al. 20 | To evaluate characteristics such as health status and specific health problems of non-participants in population-based study, and the potential resulting bias in association measures | Secondary data | Survival model, simulation | Nonparticipants were twice as likely to receive disability pensions (outcome) than participants, and even more if the pension was received for mental disorders. Simulation excluding participants with a similar profile to non-participants reduced the association between common mental disorders and the outcome |
| 18 | Manjer et al. 30 | To investigate the effect of nonparticipation on cancer incidence and mortality | Secondary data, mailed health survey | Survival model | Non-participants presented lower cancer incidence prior to recruitment and higher cancer incidence during recruitment. The proportion of participants in the cohort reporting better health was higher than in the mailed survey |
| 19 | Barchielli \& Balzi 25 | To analyze the effect on mortality of non-response in a smoking prevalence survey | Secondary data | Poisson model, life table method | All causes mortality was significantly higher among non-respondents, with higher risks for smoking related causes |
| 20 | Bergman et al. 51 | To analyze the consequences of attrition in three years after baseline in the PART study | Baseline information, sample of nonrespondents | Logistic model | Variables associated with non-participation - low income and education, non-Nordic origin and marital status - were related with depressive mood as well in the first wave |
| 21 | Petersen et al. 52 | To investigate wether terminally ill patients' reported quality-of-life scores should be adjusted for nonparticipation bias | Baseline information | Imputation methods for missing data | Significant underestimation of symptoms in 4 out of 30 comparisons suggest that imputation of quality-of-life scores of non-participants in palliative care is biased based on the available predictors |

(continues)

Table 2 (continued)

| Id | Reference | Objectives * | Data source | Analysis | Results |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 22 | Rao et al. 53 | To propose a method based on propensity scores to analytically reduce bias due to non-response | Secondary data | Propensity score based on baseline information and data imputation | Among the respondents, there was a higher frequency of women, Caucasian, married and younger people. Differences due to the proposed weighting scheme were small |
| 23 | Haring et al. 54 | To determine attrition predictors and evaluate the effect of extensive recruitment procedures on attrition and bias | Baseline information | Logistic model | The main predictors for attrition were late recruitment at baseline, unemployment, low educational level, female gender, and smoking. However attrition bias was not associated with health-related indicators |
| 24 | Van Loon et al. 31 | To investigate possible response bias in prevalence estimation and association measures | Baseline information | Logistic model | Respondents, as compared to non-respondents, presented higher socioeconomic status, better subjective health and healthier behaviors. The association measures were similar in respondents and the entire population source |
| 25 | Drivsholm et al. 55 | To compare participants at the 20-year follow-up study with nonparticipants, and to investigate the representativeness of both groups in relation to the population source | Secondary data | Logistic model | Participation decreased to $65 \%$ in the 20th follow-up year, when non-participants had lower socioeconomic status, worse health profile and higher mortality rate than participants |
| 26 | Jackson et al. 16 | To compare participants with complete clinical examinations to those with just home interview in the the ARIC study | Baseline information | Frequencies comparison | Response rates was similar for white participants, both male and female, and in all study centers. In general, respondents presented higher socioeconomic status and health, but differences were smaller for women |
| 27 | Veenstra et al. 29 | To assess association between health status at baseline and nonresponse; to analyze survival in a 5 -year follow-up | Secondary data | Logistic model | Among respondents, prevalence of coronary heart disease was higher. However, their mortality was lower than noncontacts |
| 28 | Young et al. 56 | To describe factors associated with attrition in a longitudinal study with three age cohorts of women | Baseline information | Logistic model | Variables associated with loss-to-follow-up were: education (lower), non-English-speaking origin, current smoker, poorer health and difficulty managing their income, varying according to cohort age |
| 29 | Caetano et al. 57 | To identify characteristics of nonrespondents in a survey among couples on violence and drinking | Secondary data | Logistic model | Male non-respondents were younger, less educated, more often unemployed and drinkers. Among women, having been an abuse victim during childhood increased response |
| 30 | Garcia et al. 28 | To evaluate attrition in a Spanish population-based cohort | Baseline information | Logistic model | Death and moving to another town were the main reasons of nonresponse. Refusals were associated with working status (disabled and retired) and place of birth (other regions of Spain or in foreign countries); emigration with civil status, age and education as well |
| 31 | Hara et al. 58 | To examine factors influencing the recruitment in a study collecting genetic data | Baseline information | Logistic model | Sex (male) and age (younger) presented lower participation rates. The survey location (easy access to participants' residence) and reminders sent to subjects significantly improved the participation rate |
| 32 |  <br> Thoning ${ }^{33}$ | To analyze trends in nonresponse and assess bias on morbidity prevalence | Secondary data | Logistic model | Refusals increased $4.3 \%$ in seven years (from 1987 to 1994). Nonrespondents were defined by a combination of sociodemographic characteristics. <br> Nonrespondents hospital admission rates were higher than respondents six months before data collection, and similar afterwards |

(continues)

Table 2 (continued)

| Id | Reference | Objectives * | Data source | Analysis | Results |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 33 | Jacobsen et al. 32 | To evaluate associations between socioeconomic factors and participation in the Danish National Birth Cohort | Secondary data | Frequencies comparison | Groups with low socioeconomic status were underrepresented as compared to the background population |
| 34 | Montgomery et al. 59 | To investigated potential bias due to non-participation in the follow-up of a large cohort study on pesticide applicators | Secondary data | Logistic model | Non-respondents at follow-up were younger, less educated, with lower body mass index and poorer health behaviors but better health conditions, and lower pesticide use. Estimates of exposure-disease associations did not present strong bias |
| 35 | Jousilahti et al. 60 | To evaluate total and cause specific mortality comparing participants cohort study | Secondary data | Survival model | At eight year follow up, mortality of non-participating men and women was higher than participating, except for smoking related causes |
| 36 | May et al. 61 | To evaluate potential predictors of non-response that are available at baseline (socio-economicdemographic, health, )follow-up duration and contact strategies | Baseline information | Logistic model | Age (younger), sex (male), marital status (single), poorer health conditions, and undernourishment or obesity were associated with non-response |
| 37 | Batty \& Gale 62 | To investigated variables associated with non-response and its impact on the association measures of several known risk factors and cardiovascular mortality | Secondary data | Survival model | The non-participants had higher CVD mortality than participants. However, the association measures between the risk factors evaluated and the mortality was not affected by non-response |
| 38 | Dugue et al. 63 | To estimate excess mortality comparing participants and nonparticipants in cervical screening | Secondary data | Survival model | All cause mortality and HPV-related mortality was higher for non-participants in cervical screening, and the hazard ratio increased over time |
| 39 | Hara et al. 23 | To evaluate the healthy volunteer effect comparing mortality rates among respondents and nonrespondents | Secondary data | Poisson model | Mortality was higher among nonrespondents for all causes studied, although with different effects according do sex. The relative risk varied as well according to the length of follow-up |
| 40 | Benfante et al. 64 | To investigate differences between participants and nonparticipants and the potential introduction of bias in the association measures | Secondary data | Frequencies comparison | Total mortality, cancer mortality, and coronary heart disease incidence rates were higher in non-examined men, but the differences decreased over time. No bias was found |
| 41 | Ferrie et al. 24 | To evaluate association between nonresponse at baseline and missing follow-up contacts and general mortality, and mortality by socioeconomic position | Secondary data | Survival model | Non-response at baseline and at any follow-up contact was associated with doubling the mortality hazard |
| 42 | François et al. 65 | To demonstrate how it is possible to obtain a satisfactory rate of participation in a cohort study, and to compare participants and nonparticipants | Baseline information | Frequencies comparison | The main factors associated with the response rate were: linguistic region, age, income, civil status, educational and alcohol/drugs consumption |
| 43 | Walker et al. 22 | To compare the mortality rates and the demographic characteristics between participants and nonparticipants | Secondary data | Frequencies comparison | Non-participants were younger, more likely to be unmarried and work in less skilled jobs. Their mortality rates were higher in the first three years of follow-up, decreasing afterwards. CVD mortality was similar in both groups |
| 44 | David et al. 66 | To assess the performance of two different models with two end points each, in analyzing loss-to-follow-up | Secondary data | Logistic and survival model | Survival models performed better than logistic models |

[^1]Table 2 (continued)

| Id | Reference | Objectives * | Data source | Analysis | Results |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 45 | Froom et al. 67 | To investigate the healthy volunteer effect in an occupationally cohort of male industrial employees | Secondary data | Survival model | All cause mortality hazard ratio was higher in nonparticipants, and the difference persisted up to 8 years of follow-up |
| 46 | Bopp et al. 68 | To evaluate feasibility and quality of linkage procedure in providing followup information | Secondary data | Survival model | Linkage success was independent of any variables. Losses in 10 years were $4.7 \%$. Participants of the study had lower mortality than the general population |
| 47 | Criqui et al. 69 | To evaluate differences in cardiovascular health status according to participation in a population based study | Baseline information, non-respondents telephone interview | Frequencies comparison | Non-respondents presented more CVD but did not differ on known hypertension. Impact on prevalence estimates was small due to low proportion of nonresponse |
| 48 | Lindsted et al. 70 | To assess the healthy volunteer effect comparing mortality rates between the respondents to a small questionnaire with respondents to a full detailed questionnaire | Secondary data | Survival model | Hazard ratio for different mortality causes was larger for non-respondents, but the difference decreased over time |
| 49 | Thygesen et al. 71 | To estimate the effect of drop-out on the association between alcohol intake and mortality | Secondary data | Poisson model | Loss to-follow-up was associated with increased mortality and incidence rates of heart disease, some cancers, and liver diseases related to alcohol intake |
| 50 |  <br> Rasmussen 72 | To evaluate if baseline characteristics could provide sufficient information about non-response bias | Secondary data | Logistic model | At baseline, respondents and non-respondents presented similar profiles (smoking, lung function and respiratory symptoms). However, non-respondents had larger rates of hospital admission due to respiratory diseases, indicating that equal baseline profile does not protect against non-response bias |

CVD: cardiovascular diseases.

* Objectives presented here were the most related to the objective of this review.
ated with participation, in spite of its importance in some of them $11,18,19,20,21,22,23,24,25$.

Study region showed no association with participation, in spite of the diversity of places evaluated. Participating in studies voluntarily, giving time, information and biological material is all related to ideas of social capital and volunteering ${ }^{16}$, and we expected variation according to local cultural components.

Participation in studies has also been associated with behavioral variables and with general state of health. Non-participants report greater consumption of alcohol, smoking and poor general state of health $12,15,2019,21,22,23,26,27,28,29,30,31,32$, 33,34. This information, however, are not available in most publications, limiting the scope of our study.

Strategies to increase participation proportion have been proposed in terms of persuading individuals who are reluctant or hesitant; however, willingness to participate is not always accompanied by commitment to adhere to the study in the long term ${ }^{35}$. Lastly, we agree with the argument of

Morton et al. 3 that more information should be requested on the profile of participation and its potential bias.

There is a major need to pursue methodological studies to evaluate the impacts of non-participation on measures of effect in cohort studies. Strategies for that kind of evaluation include comparing participants with non-participants through administrative data bases (sex, age, place of residence), application of summary questionnaires and passive follow-up of eligible population to evaluate mortality ${ }^{4}$. Recent publications from journals with high impact factors show that nonparticipation is mostly ignored or dismissed by many authors, although some are attempting to reduce it or mention it as a limitation in their study ${ }^{36}$.

In conclusion, our findings suggest that the drive for participation and compliance should be assessed previously to funding the cohort study, and specific local knowledge should be included in addressing the potential participants.

Figure 2
Correlation of year the baseline year and participation rate.


Table 3

Univariate and multiple meta-regression models.

| Variables | VAF | Univariate meta-regression models |  |  | Multiple meta-regression modelsVAF = 18.1\% |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |
|  |  | OR | 95\%CI | $p$-value | OR | 95\%CI | p -value |
| Age (mean) | 2.6\% | 0.99 | 0.96-1.01 | 0.24 | 0.97 * | 0.95-1.00 | 0.04 |
| Proportion of women | 6.7\% | 1.01 | 0.99-1.03 | 0.21 | 1,01 | 0.99-1.03 | 0.18 |
| Baseline year | 11.9\% | 0.97 | 0.95-0.99 | 0.01 | 0.97 * | 0.94-0.99 | 0.02 |
| Selection (baseline: Sampling) |  |  |  |  |  |  |  |
| Census | 0.0\% | 0.99 | 0.66-1.48 | 0.95 | - | - | - |
| Population (baseline: General |  |  |  |  |  |  |  |
| Other | 0.0\% | 1.02 | 0.63-1.67 | 0.93 | - | - | - |
| Outcomes (baseline: General health) |  |  |  |  |  |  |  |
| Cardiovascular diseases | 3.6\% | 0.78 | 0.49-1.24 | 0.30 | - | - | - |
| Other |  | 1.12 | 0.70-1.80 | 0.65 | - | - | - |
| Study region (baseline: Northern Europe) |  |  |  |  |  |  |  |
| U.S.A. | 7.5\% | 1.17 | 0.67-2.04 | 0.59 | 0.94 | 0.51-1.73 | 0.85 |
| Continental Europe |  | 0.69 | 0.43-1.12 | 0.13 | 0.64 | 0.38-1.07 | 0.09 |
| Other |  | 1.10 | 0.65-1.89 | 0.72 | 0.94 | 0.52-1.69 | 0.83 |

$95 \% \mathrm{Cl}: 95 \%$ confidence interval; OR: odds ratio; VAF: variance accounted for.

* For the change of one unit in the variable causes decline the odds of participation.


## Resumen

La proporción de no participación en estudios de cohorte se asocia también con la exposición y probabilidad de ocurrencia de hechos que pueden generar sesgos en las estimaciones de interés. El objetivo de este estudio es realizar una revisión sistemática y un metaanálisis de artículos que describen la participación en estudios de cohortes y evaluar las características asociadas con la participación. Una revisión sistemática fue realizada (MEDLINE, Scopus y Web of Science), en busca de artículos que describen la relación de participación basada en estudios de cohortes. Se seleccionaron 2964 artículos, de los cuales se identificaron preliminarmente 50. Entre estos, la proporción promedio de participación fue de un 64,7\%. Utilizando la metarregresión, sólo la edad, años de referencia y la región de estudio (borderline) se asociaron con la participación. Teniendo en cuenta la disminución de la participación en los últimos años, y el coste de los estudios de cohortes, es esencial buscar información para evaluar el potencial de la no participación antes de comprometer recursos.

[^2]
## Contributors

S. H. A. Silva Junior was responsible for the first draft of the manuscript and data analyzes and contributed to the conception and design of the study. S. M. Santos, C. M. Coeli and M. S. Carvalho contributed to the conception and design of the study. All authors contributed significantly to interpreting the results, commented extensively on subsequent revisions, and read and approved the final manuscript.

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[^0]:    (continues)

[^1]:    (continues)

[^2]:    Sesgo de Selección; Estudios de Cohortes; Métodos Epidemiológicos

